Insider Information and Performance Pay

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Abstract

This paper provides evidence that managers have private information they exploit for financial gain at the expense of shareholders. It develops a model of optimal contracting to show that moral hazard, hidden actions taken by agents, can rationalize why a principal would optimally induce agents to benefit from their private information. Estimates from a structural model shows that moral hazard is an important economic factor. This leads to the conclusion that, in practice, shareholders and managers might optimally agree upon an arrangement where managers systematically exploit their private information about the firm.

I. Introduction

This paper provides evidence that managers exploit information hidden from shareholders for personal gain through their compensation contracts. We demonstrate this can arise as a part of an optimal arrangement between shareholders and managers, in a model where managers' actions are also hidden from shareholders and their board of directors, creating a moral hazard problem. We estimate a structural model of moral hazard that quantifies its cost, and find that moral hazard from hidden actions is an important determinant of managerial compensation packages. So from a public policy perspective, one should not interpret evidence that managers benefit from their insider information as indicative of failure in corporate governance. Optimal compensation plans may induce managers to benefit from their private information, even if shareholders do not face public choice or free rider problems when aggregating their interests, and can easily communicate their wealth maximizing goal to managers through their board of directors.

The paper is organized around three questions. Are managerial compensation packages more likely to be tilted towards performance pay, such as dividends and capital gains from stock and options in the firm, when their firms subsequently perform well on the stock market? Can theory rationalize why shareholders would sanction this practice under some circumstances? The third question asks whether there are high economic benefits from shareholders creating incentives that align the manager's interests to theirs.

Based on evidence we provide in Section 2 from a longitudinal data set which tracks about 1500 U.S. firms and their managers for nine years, we answer the first question in the affirmative. Changes in the stock components of the manager's compensation package are a significant variable in forecasting financial returns. One might interpret this empirical finding as a trade-off between higher net financial returns to shareholders and more risky financial packages for managers, but this implies some firms do not maximize net financial returns. We are reluctant to drop the neoclassical premise that shareholders maximize the expected value of their firm. We would rather interpret the affirmative answer to the first

question as evidence that managers have insider information, which helps them decide how much of their compensation package should be fixed and how much should be affected by the firm's performance.

Our preferred interpretation begs the second question. In Section 3 we develop a principal agent model in which moral hazard (or hidden actions taken by the agent) is a necessary condition to rationalize why the principal induces the agent to benefit from their private information about the state of the firm.

Unless moral hazard is an important empirical issue, it is hard to rationalize why shareholders would tolerate executives who benefit from private information about the firms they manage. Thus in Section 4 we answer the third question by estimating a structural model of moral hazard and find that the economic costs of ignoring moral hazard are indeed substantial.

There are many reasons why firm executives are probably much better informed than shareholders about the prospects of the enterprise, and its demands on managerial time, energy and expertise. As opportunities to make the firm more profitable are explored, management gain foresight into which ventures are likely to be successful, and those which will probably fail, putting them in a favorable position to trade on their insider knowledge. If a manager could choose how many firm specific assets to hold without incurring penalties directed at those who engage in insider trading, he might prefer holding more stock and options in his own firm when his private prognosis was more favorable than the market's, and less firm specific assets when his insider knowledge projects a worse outcome than what stockholders and other investors think.

Insider trading is illegal, and the Securities and Exchange Commission is charged with combatting insider profits from arbitrage.[1] But the SEC does not prevent a manager from

[1] Harris (2003) describes how the SEC prepares to prosecute cases of alleged insider trading. Large volume transactions accompanied by big price shifts are a signal that information about the firm's prospects may have been exploited by insiders. When alerted to a possible infringement (perhaps by a trader who believes he was exploited by an insider), the SEC compiles a list of investors who traded during the period under consideration, the insiders privy to information that led to the price change, and tries to match parties from both lists.

benefiting from insider information by holding more of his company's stocks when he believes his firm's medium term prospects look more favorable than usual. Shareholders could prevent a manager from benefiting from his detailed knowledge of his firm. Existing regulations in the United States require the manager to frequently report all trading in the firm's assets. Hence it would be a relatively straightforward for the board of directors to penalize the manager, and thus prevent him, from ever holding any of the firm's assets. In the absence of moral hazard and/or the opportunity to benefit from inside trading, it is hard to imagine why a manager would prefer to hold financial assets in his own firm compared to the alternative of holding a well diversified portfolio. Managers of nonprofit enterprises and high ranking government officials are routinely required to divest themselves of assets that may cause a conflict of interest between their professional role and personal wealth management. Therefore managers should have no more qualms about agreeing to such a requirement, than agreeing to rules governing company perks, or theft of company property. Thus the board could greatly curb if not entirely eliminate the benefits managers garner for themselves from their inside information by issuing an easily enforced directive to their executive management.

In practice this never occurs. A large portion of executive compensation is tied to firm specific assets. From an empirical standpoint, trading by corporate insiders appears to be profitable. Seyhun (1986) finds that insiders tend to buy before an abnormal rise in stock prices and sell before an abnormal decline. Earlier studies by Lorie and Niederhoffer (1968), Jaffe (1974), and Finnerty (1976) draw similar conclusions. More recently, Seyhun (1992a) finds compelling evidence that insider trading volume, frequency, and profitability all increases significantly during the 1980s. Over the decade, he documents that insiders earned over 5 percent abnormal returns on average. Seyhun (1992b) determines that insider trades predict up to 60 percent of the total variation in one-year-ahead returns. To summarize, hidden information is an economically important phenomenon in executive compensation.

So it is paradoxical that managers are compensated on the basis of their firm's perfor-

mance, such as dividends and capital gains, when the profitability of the firm partly depends on how managers assess their own accomplishments and firm's prospects. Bebchuk and Fried (2003) and others have argued that one reason why managers are paid stock options instead of assets that are easier to value, such as cash, is that shareholders systematically underestimate option expenses. Similarly Bertrand and Mullainthan (2000, 2001) argue that the separation of management from ownership in public corporations allows the CEO to gain effective control of the pay-setting process. They argue that skimming is less likely to attract the attention of shareholders when the firm performs well. Consequently granting options should be an excellent vehicle for skimming, costing shareholders nothing when the firm performs poorly.

Such explanations are most convincing when there are costs of governance at the level of the board of directors from administering shareholder goals of maximizing financial returns. However, the fact that managers receive restricted stock and options, which tie their personal wealth to the fortunes of their employers, does not signify governance problems at the board level. In principal agent models with hidden actions but no private information, shareholders could use these instruments to enforce the optimal contract. Margiotta and Miller (2000) develop econometric methods for estimating a structural model of moral hazard and find it is a very important factor for explaining variation in managerial compensation. Moreover using similar estimation methods, we find in Gayle and Miller (2008a), that the structure of compensation, conditional on firm characteristics, has been stable over the last 60 years. That paper shows the secular increase in the mean and variance of managerial compensation is largely explained by the changing distribution of firm size occurring throughout this extended period. Both studies are applied to relatively small data sets taken from only three industries, namely aerospace, chemicals and electronics. The empirical work on the costs of moral hazard reported in Section 4 extends these other results to a broad industry spectrum of publicly traded firms.

II. Insider Wealth, Abnormal Returns and Compensation

This section addresses the first question raised in the introduction. After briefly describing the longitudinal data set compiled for undertaking the empirical work, we conduct a regression analysis of the empirical evidence on the benefits of insider information. We focus on changes in stockholding that occur before the period begins to see whether they help predict future returns. Using a model with a simple linear decision rule for insider trading, we test whether managers condition on more information than the market does in forming their expectations about future returns. Regressing the manager's portfolio choices on next period's abnormal returns to the firm, we find the latter are positive and significant, evidence that future returns are a noisy indicator of inside information available to the manager.

We then conduct simulations in an attempt to quantify the magnitude of the gains to managers from their insider information. We construct a simple dynamic portfolio strategy based on changes in asset holdings by managers, and find that this strategy significantly outperforms the market.

In the latter parts of Section 2 we investigate whether, conditional on the information held by the manager, compensation to managers fluctuates with the idiosyncratic components of the return to his firm. After controlling for the manager's portfolio choices and other observed factors that affect abnormal returns, we find that unexplained variation in abnormal returns is positive and significant. This result supports the hypothesis that rather than simply benefiting from their private information by structuring their compensation package in an advantageous way, managers are also motivated, through their work choices, to raise the mean of unanticipated abnormal returns. These new findings suggest a second explanation for why shareholders do not prevent managers from personally exploiting their insider knowledge about the firm. If their actions are also hidden and affect firm performance, then not linking the manager's wealth to the firm's value might create a moral hazard problem, an issue we take up in the sections which follow.

A. Data

Our analysis is based on longitudinal data gathered from three main sources: Standard & Poor's ExecuComp, Compustat databases, and Executive Compensation Reports data on firm compensation plan responses to Section 162(m). Our database tracks about 1,500 firms over a nine year panel beginning in 1992 in the S&P 500, Midcap, and Smallcap indices. It contains information on at least the five highest paid executives for 1,837 unique CUSIP identifiers.[2] For much of our work we partitioned firms by the ten sectors labeled as Energy, Materials, Industrials, Consumer Discretionary, Consumer Staples, Health Care, Financial, Information Technology, Telecommunications Services, and Utilities

Table 1

Summary Data on Firms: 1992-2002

(Sales, Equity, and Assets are in millions of US$ (2000))

Variable	Mean	Standard Deviation
Abnormal Returns	0.024	0.431
Return on Assets	1.42	25.96
Sales	3023	6754
Total Equity	1316	3199
Total Assets	3054	6686
No. of Firms	1517	
No. of Observations	12470	

Our data is similar to those used in Gayle and Miller (2008a) to investigate how managerial compensation has changed over the last sixty years.[3] Table 1 shows the average

[2]The SEC mandates that firms must submit compensation data on the top five executives, but many firms report data on a greater number. We used data on the top six when it was available, and the top five otherwise.

[3]The variables used in this analysis are defined as in Gayle and Miller (2008a), where details are provided about their construction. Gayle and Millller (2008a) use three samples in their empirical analysis The first tracks selected firms in the aerospace, chemicals and electronics industries up until 1977. The third is compiled

firm size using three measures, sales, equity and assets. The standard deviations are about twice to three times as large as the sample means. We report two measures of income, the return on assets and abnormal returns. The latter are defined for the n^{th} firm at time t as $x_{nt} = \pi_{nt} - \pi_t$, where π_t denotes the return on the market portfolio in period t and π_{nt} is the firm's financial return. Thus x_{nt} is a relative measure that uses stock market performance as a benchmark. We also experimented with other benchmark performance measures such as industry and sector returns, but they do not significantly affect the results reported below. Note that measures of income are much more dispersed than the measures of firm size with standard deviations about twenty times the respective sample means.

Total compensation and one of its components, pretax salary and bonus, are summarized in Table 2. The other main components in executive compensation are stock and option grants, vested retirement benefits, as well as gains and losses from abnormal returns on stocks and other financial securities in the manager's portfolio. The reason for including the last component is that outsiders eliminate firm specific risk by holding only a negligible amount of any given firm's securities from their wealth portfolios, in this way guaranteeing the return on the market portfolio, rather a random variable distributed about that return. Thus the fact each manager is so heavily vested in his own firm indicates a professional interest that comes with his job. We report averages for the CEO, as well as the next five highest paid executives in the firm, along with the respective standard deviations.

in exactly the same way as the data in this paper, but for a longer time period, namely 1992 through 2003 (as opposed to 2001), and for fewer executives only (the top 3 as opposed to the top 6). The second is a subset of the third, selecting only those observations from the third sample in the three industries mentioned above. Since the sample used in this paper is very similar to the third sample, we refer readers to a more detailed summary of the third sample provided in Tables 1 though 4 of our companion paper.

Table 2

Summary Data on Executive Compensation in Millions of US$ (2000)

Variable	Rank	Mean	Standard Deviation
Pretax-tax Compensation			
	All	2.01	13.44
	CEO	2.31	18.46
	Non-CEO	1.96	12.47
Pretax Salary and Bonus			
	All	0.699	0.656
	CEO	0.919	1.07
	Non-CEO	0.655	0.528
No. of Observations			
	All	81578	
	CEO	12460	
	Non-CEO	69108	

The main patterns in this data set are reflected in many other samples of executive compensation. The CEO receives less than half of his compensation in salary and bonus, which exhibits much lower variability than the sum of the other components, both between and within industries or sectors. Lower ranked officers are paid less than CEOs and receive a higher proportion of their pay in salary and bonus.

B. Executive portfolio choices and future returns

If managers were more informed than the market and were able to exploit this information for personal gain, it seems reasonable to conjecture that the information impounded in future returns would help predict what choices managers had already taken with respect to their wealth portfolios. In that case abnormal returns would be a noisy predictor of ret-

rospective choices. We now denote with a t subscript information available to stockholders at time t, and denote with a $t+\Delta$ subscript information available to the managers at time t. We define $u_{n,t+\Delta}$, the conditional expectation of the abnormal return at the beginning of period $t+1$ based on all the information available to the manager in period t, as:

$$u_{n,t+\Delta} \equiv E_{t+\Delta}\left[x_{n,t+1}\right]$$

and denote by $q_{n,t+\Delta}$ stock purchases by the manager in period t. Defining:

$$\varepsilon_{n,t+1} \equiv x_{n,t+1} - u_{n,t+\Delta}$$

as the manager's forecast error of abnormal returns and (in this section only) assuming that the manager's decision rule for trading is linear in $u_{n,t+\Delta}$, it follows that:

$$q_{n,t+\Delta} = \delta_0 + \delta_1 u_{n,t+\Delta} \equiv \delta_0 + \delta_1 x_{n,t+1} + \delta_1 \varepsilon_{n,t+1}$$

where, by the definition of $u_{n,t+\Delta}$:

$$E_{t+1}\left[\varepsilon_{n,t+1}\left|u_{n,t+\Delta}\right.\right] = 0$$

If we impose the additional restriction that $\delta_0 = 0$, then this decision rule may be interpreted as a linear approximation to the optimal rule for a risk averse expected utility maximizer confronted with a favorable gamble. When $\delta_0 = 0$ the rule implies that $q_{n,t+\Delta} \equiv 0$ if and only if $u_{n,t+\Delta} \equiv 0$. From the definition of $u_{n,t+\Delta}$, this is true if and only if $E_{t+\Delta}\left[x_{n,t+1}\right]$ and the unconditional expectation, $E_t\left[x_{n,t+1}\right] = 0$, are the same. In that case insider trading is conducted if and only if the manager has insider information about next period's abnormal

return. Regressing $q_{n,t+\Delta}$ on $x_{n,t+1}$ we obtain a consistent estimator of:

$$\frac{E\left[q_{n,t+\Delta}x_{n,t+1}\right]}{E\left[u_{n,t+1}x_{n,t+1}\right]} = \delta_1\left(1 + E\left[\varepsilon^2_{n,t+1}\right]\right)$$

The expression is positive if and only if $\delta_1 > 0$.

The results from running this regression are reported in Table 3. The coefficient on lead abnormal return, δ_1, is positive and significant in the sample as predicted by this simple model of insider information. Also consistent with the simple linear model δ_0, the constant term, is insignificant. In the same regression we also included the ratio of (contemporaneous) salary and bonus to total compensation to investigate whether the manager takes a lower salary and bonus in return for more claims that are contingent on the firms' return. Although the sign of δ_2 is negative, it is not statistically significant. The lack of significance should not, however be interpreted as evidence against the model, since the manager is free to draw from his own outside wealth to invest in his firms' stock when promising prospects arise.

Table 3

OLS Estimates of the Effect of Future Return on Changes in Managers Stock Holdings.

Dependent Variable: Yearly Change in Number of Shares Managers Hold in their Firms

Variable	Estimate
Ratio of Salary and Bonus to Total Compensation	−0.77
	(2.13)
Lead Abnormal Return	2.30
	(1.11)
Constant	80.34
	(50.21)
R^2	0.12
No. of Observations	67769

Note: Estimated standard errors in parentheses.

Managers are required to report all their trading activity to the SEC within a month, and their reports are available for public scrutiny. Consequently our finding that managers appear to exploit inside information when investing in their own firm raises the possibility that others might be able to benefit from their serendipitous choices. Table 4 presents our findings from regressing abnormal returns on the manager's lagged trading activity, providing some evidence of how well their trading activity is a useful predictor of abnormal returns. (Dummy variables for the sectors were also included in the regression, but are not reported here.) The estimated coefficients in question are positive and significant in both regressions, consistent with the hypothesis that managers exploit insider information. The estimates also show there is a negative relationship between abnormal returns of the firm and the ratio of salary and bonus to total compensation, but again the relationship is statistically insignificant, reinforcing our earlier point that resources used for insider trading need not come at the expense of other components in the compensation package, but could simply reflect an adjustment in the manager's asset portfolio.

Table 4

OLS Estimates of the Effect of Lagged Change in Manager's Stock Holdings on Firms Abnormal Returns. Dependent Variable: Yearly Real Abnormal Return

Regressors	Estimate
Lagged Change in Manager's Stock Holdings	$2.9e-04$
	$(8.0e-05)$
Ratio of Salary and bonus to Total Compensation	$-8.2e-03$
	$(5.2e-01)$
Lagged Return on Book Value of Assets	$-4.1e-03$
	$(4.7e-04)$
Lagged Dividends per Share	$-3.5e-02$
	$(9.5e-03)$
Lagged Return on Market Value of Equity	$-4.2e-04$
	$(5.9e-05)$
Lagged Earnings per Share	$3.75e-06$
	$(1.2e-04)$
R^2	0.04
No. of Observations	67769

Note: Estimated standard errors in parentheses.

Much of the evidence from Tables 3 and 4 supports the notion that managers exploit their superior knowledge about their own firm's performance on the stock market, but not all. As above, suppose the manager follows the linear decision rule for insider trading, and has access to the other regressors listed in Table 4, which we now call z_{nt}. In this case the inverse of the coefficient on lagged changes in the manager's stock holdings is δ_1, and the coefficients values on all the other variables are zero because:

$$x_{n,t+1} = \delta_1^{-1} q_{n,t+\Delta} - \delta_0 \delta_1^{-1} - \varepsilon_{n,t+1}$$

and the manager's forecast error satisfies the conditional expectation $E[\varepsilon_{n,t+1}|z_{nt}] = 0$. Hence, our finding that several coefficients are significant, constitutes evidence against the linear model. We also note that an estimate of $E\left[\varepsilon^2_{n,t+1}\right]$ can obtained by subtracting 1 from the product of the estimated coefficient on $x_{n,t+1}$ in Table 3 and the estimated coefficient on $q_{n,t+\Delta}$ in Table 4. The estimated variance is negative, casting further doubt on the linear specification.

C. Gains from insider trading

To gauge the magnitude of the gains from insider trading, we conducted a simulation exercise to retrospectively evaluate how lucrative it would have been to base a portfolio investment strategy on data from these reports over the 9 year period covered by the new data set. The simulations generated the outcomes of three self financing strategies. The first strategy is an outsider strategy, to invest in the market portfolio. The third strategy is only feasible if the inside investor perfectly anticipates the one period ahead abnormal return of the companies; an investor privy to perfect inside information pertaining to the n^{th} firm invests all his wealth in its shares in period t if $\pi_{n,t+1} > \pi_{t+1}$ and all of it in the market portfolio if $\pi_{n,t+1} \leq \pi_{t+1}$, reaping a certain return for the period of:

$$\pi^{(3)}_{n,t+1} \equiv \max\{\pi_{n,t+1}, \pi_{t+1}\}$$

Note there is an upper bound to the gains from perfect foresight because it is self financing strategy after the initial outlay.

The second strategy allocates a fraction of the manager's discretionary wealth, λ_{nt}, to the market portfolio in period t, and the remaining proportion $(1 - \lambda_{nt})$ to stock in the n^{th} firm for a return of:

$$\pi^{(2)}_{n,t+1} = \lambda_{nt}\pi_{t+1} + (1 - \lambda_{nt})\pi_{n,t+1}$$

where λ_{nt} reflect the historical portfolio choices of the n^{th} manager as observed in the data.

Here discretionary wealth is defined as the difference between the maximum observed wealth observed the executive in the firm observed over the sample period, denoted by $\overline{W}_n$, and the minimum, denoted by $\underline{W}_n$. Thus λ_{nt} is defined by:

$$\lambda_{nt} \equiv \frac{W_{nt} - \underline{W}_n}{\overline{W}_n - \underline{W}_n}$$

We compared the outcomes of these three investment strategies, to see whether following the reports managers submit would have been profitable, and how much of the potential gains from clairvoyance managers are able to extract. The market return averaged almost 1.089 per year in this period (with standard deviation 0.097), but if an executive could have perfectly anticipated returns in his own firm, this number increases to 1.192 (with standard deviation 0.268). More surprising is our result that almost all these gains are realized by following the second strategy we defined, which produced an average annual return of 1.196 (with standard deviation 0.336). Note the standard arguments for diversification imply that replicating the second portfolio strategy in each firm would not have exposed investors to any idiosyncratic risk.

As a final check we investigated whether the cumulative gain from following these different strategies are statistically significant from each other, by testing the null hypothesis:

$$\lim_{N\to\infty} \frac{1}{N} \sum_{n=1}^{N} \left[\prod_{t=1}^{T} \left(\pi_{nt}^{(i)} \right) - \prod_{t=1}^{T} \left(\pi_{nt}^{(j)} \right) \right] = 0$$

for various $(i, j) \in \{1, 2, 3\}$ where $\pi_{nt}^{(1)} \equiv \pi_t$ is the market return and $n \in \{1, \ldots, N\}$ is the sample population of executives. The null hypothesis is strongly rejected for all the various combinations. To summarize, while perfect foresight beats everything, building an investment strategy based on the manager's stock holding is also significantly more profitable than specializing in the market portfolio.

D. Evidence for moral hazard

The evidence presented above favors the view that managers undertake insider trading, exploiting privy information to trade in their firm's stock at the expense of shareholders. We argued in the introduction that these activities are tacitly or explicitly approved by their respective boards of directors because insider trading by managers could be greatly curbed or even eliminated. Boards could require managers to refrain from owning financial assets of the firms they manage. After all certain positions in the public sector, such as elected officers, require the occupant to divest himself of assets in firms that might create a conflict of interest between his professional duties and the incentives of the firms' shareholders. One reason why boards might be reluctant to discourage insider trading is that compensation from insider trading might help align incentives between shareholders and the manager. If so, executive compensation packages might also depend on those components of abnormal returns that are not anticipated by inside knowledge.

Recall $x_{n,t+1}$ is the abnormal return in the upcoming period $t+1$ and $\varepsilon_{n,t+1}$ is the residual of abnormal returns that the manager of firm n does not anticipate. Let $w_{n,t+1}$ denote his compensation paid at the beginning of the next period $t+1$. If insider trading does not fully resolve the conflicts of interest between shareholder and management objectives, then the board of directors should make $w_{n,t+1}$ depend positively on $\varepsilon_{n,t+1}$.

Since $\varepsilon_{n,t+1}$ is unobserved, we regressed $w_{n,t+1}$ on a estimate of $\varepsilon_{n,t+1}$, simultaneously controlling for other variables that managers use in forming their expectations about $x_{n,t+1}$. Based on the identity:

$$\varepsilon_{n,t+1} \equiv x_{n,t+1} - u_{n,t+\Delta}$$

and recalling $u_{n,t+\Delta} \equiv E_{t+\Delta}\left[x_{n,t+1}\right]$ is his conditional expectation in period t about $x_{n,t+1}$, we formed:

$$\widehat{\varepsilon}_{n,t+1} \equiv x_{n,t+1} - \widehat{u}_{n,t+\Delta}$$

from the estimated expectation function presented in Table 4. Then we regressed $w_{n,t+1}$ on

$\widehat{\varepsilon}_{n,t+1}$ as well as the variables used in estimating $\widehat{u}_{n,t+\Delta}$. Our estimates in Table 5 show that managers are rewarded (punished) when the unanticipated component of abnormal returns is higher (lower) than they expected. This suggests that shareholders are not only less informed about the economic prospects of their firm, but also that shareholders do not fully monitor the activities of their management.

Table 5

OLS Estimates of the Effect of Unanticipated Changes in Abnormal Return on Total Compensation. Dependent Variable: Yearly Total Income Compensation in Thousands of US$ (2000)

Regressors	Estimate
Unanticipated change in abnormal return	726
	(89)
lagged change in stock holdings	11.6
	(1.2)
lagged return on assets	16.9
	(6.7)
lagged dividends per share	60.4
	(115)
lagged return on equity	0.56
	(0.85)
lagged earnings per share	-0.95
	(1.65)
constant	2504
	(98)
R^2	0.28
No. of Observations	67769

Note: Estimated standard errors in parentheses.

III. Generalized Moral Hazard

This section answers the second question using a principal agent model borrowed from Gayle and Miller (2008b) that analyzes interactions between insider information, moral hazard and executive compensation. We take up the idea that both insider information and moral hazard might play a role in contracting with managers. In our model shareholders do not observe the manager's activities. After paying the manager for his work in the previous period, at the beginning of each period the board of directors proposes a compensation plan to the manager, which depends on the realization of the firms abnormal returns as well as accounting information to be later provided by the manager. Based on the board's proposal the manager decides whether to remain with the firm or leave and picks real consumption expenditure for the period. Having accepted the contract offer, the manager observes the firms prospects, provides some accounting information, and chooses a work routine that is not observed by the directors. The return on the firms assets are realized at the end of the period. It depends on how well the firm was managed during the period, the private information available to the manager, as well as other unanticipated factors.

The objective of the manager is to sequentially maximize his expected lifetime utility, and the goal of the firm is expected value maximization. An optimal contract between shareholders and executives satisfies three conditions, a participation constraint, that assures the manager he will have higher expected utility from employment with his firm rather than another one, an incentive compatibility constraint, that induces him to maximize the value of the firm rather than using the resources of the firm to pursue some other objective, and truth telling constraint that induces the manager to reveal his inside information. We show that absent moral hazard, the optimal contract is to pay the manager a fixed compensation irrespective of his private information. Furthermore, if insider information is not intrinsically linked to the moral hazard problem in a sense we make explicit, then although the optimal contract should depend on the firm's abnormal returns, permitting the manager to exploit his nsider information is suboptimal. Finally, if moral hazard and private information are

intrinsically linked, then the gains from private information can be incorporated into the optimal contract. Rather than discourage managers from benefitting from their insider knowledge, the board might optimally sanction it.

A. The model

More specifically, at the beginning of period t the manager is paid compensation denoted w_t for his work in period $t-1$ according to the schedule the shareholders had previously committed, and his managerial contract is up for renewal. He makes his consumption choice, a positive real number denoted by c_t, and the board proposes a new contract. At that time the manager chooses whether to be engaged by the firm or be engaged outside the firm, either with another firm or in retirement. Denote this decision by the indicator $l_{t0} \in \{0,1\}$, where $l_{t0} = 1$ if the manager chooses to be engaged outside the firm and $l_{t0} = 0$ if he chooses to be engaged inside the firm. If $l_{t0} = 0$, the prospects of the firm are then fully revealed to the manager but partially hidden to the shareholders.

We assume throughout that managers privately observe $s_t \in \{1,2\}$ in period t, information that affects the distribution of the firm's abnormal returns. The board announces how managerial compensation will be determined as a function of $s'_t \in \{1,2\}$, what he tells them about the firm's prospects, and its subsequent performance, as measured by abnormal returns x_{t+1} revealed at the beginning period $t+1$. The manager truthfully declares or lies about the firm's prospects by announcing $s'_t \in \{1,2\}$, effectively selecting a schedule $w(s'_t, x_{t+1})$ indexed by his announcement s'_t. He then makes his unobserved labor effort choice, denoted by $l_{tj} \in \{0,1\}$ for $j \in \{1,2\}$ in each period t. There are two possibilities, to work diligently for the firm by pursuing the shareholders objectives of value maximization, and indicated by setting $l_{t2} = 1$, or to be employed by the firm but shirk, following different objectives than maximizing the firm's value, and here denoted by $l_{t1} = 1$.

At the beginning of the period $t+1$ abnormal returns x_{t+1} for the firm are drawn from a probability distribution which depends on the true state s_t and the manager's action

l_t. We denote the probability density function for abnormal returns in period t when the manager works diligently and the state is s by $f_s(x_{t+1})$, and let $f_s(x_{t+1})\, g_s(x_{t+1})$ denote the probability density function for abnormal returns in period t when the manager shirks, bounded below by the same real number ψ. Note that $g_s(x)$ is the likelihood ratio for abnormal returns from shirking versus working diligently in state s. We assume that the firm's losses from shirking do not depend on the state, meaning $f_1(x)\, g_1(x) \equiv f_2(x)\, g_2(x)$. Consequently higher profits from being in the better state can only be realized if the manager is diligent.

Preferences over consumption and work are parameterized by a utility function exhibiting absolute risk aversion that is additively separable over periods and multiplicatively separable with respect to consumption and work activity within periods. Lifetime utility is expressed as:

$$-\sum_{t=0}^{\infty}\sum_{j=0}^{2}\beta^t\alpha_j l_{tj}\exp(-\rho c_t)$$

where β is the constant subjective discount factor, ρ is the constant absolute level of risk aversion, and α_j is a utility parameters with consumption equivalent $-\rho^{-1}\log(\alpha_j)$ that measures the distaste from working at level $j \in \{0,1,2\}$. We assume $\alpha_2 > \alpha_1$ meaning that compared to the activity called shirking, diligence is more aligned to the shareholders' interest than the manager's interests, and without loss of generality scale utility so that $\alpha_0 = 1$. This simply means that α_j values the nonpecuniary features of engaging in activity $j \in \{1,2\}$ within the firm relative to the total current utility value from leaving the firm.

The manager's wealth is endogenously determined by his consumption and compensation. We assume there are a complete set of markets for all publicly disclosed events, effectively attributing all deviations from the law of one price to the market imperfections under consideration.

B. The optimal contract

Drawing upon the work of Fudenberg, Holmstrom and Milgrom (1990), we can show that the optimal long term contract solved by shareholders can be implemented by a sequence of short term contracts. Appealing to Myerson (1982), the revelation principle implies that, without loss of generality, we can restrict the set of feasible contracts to those that respect the participation, incentive compatibility and truth telling constraints we now define. The participation constraint states that the manager prefers working one more period and then leaving to not working for the firm at all. The incentive compatibility constraint restricts short term contracts to those payment schedules in which the manager prefers to work diligently rather than shirk. The truth telling condition requires shareholders to write contracts that induce the manager to select a compensation schedule that reveals the firm's prospects. We assume throughout that legal considerations induce the manager not to overstate the firm's prospects but that incentives must be provided to persuade the manager from understating them.

This leads to the following formulation of the optimization problem shareholders solve, to maximize:

$$\begin{aligned}
&\sum_{s=1}^{2} \varphi_s \int_{\psi}^{\infty} \log v_s(x) + \eta_0 \left[\alpha_2^{1/(1-b_t)} - v_s(x)\right] f_s(x)\, dx \\
&+ \sum_{s=1}^{2} \varphi_s \eta_s \int_{\psi}^{\infty} v_s(x) \left[\left(g_s(x) - (\alpha_2/\alpha_1)^{1/(b_t-1)}\right] f_s(x)\, dx \right. \\
&+ \varphi_2 \eta_3 \int_{\psi}^{\infty} \left[v_1(x) - v_2(x)\right] f_2(x)\, dx
\end{aligned}$$

with respect to $v_s(x) \equiv \exp\left[-\rho w_s(x)/b_{t+1}\right]$, which measures how utility is scaled up by compensation if abnormal returns x are realized at the end of the current period t when state s is announced, where b_{t+1} is the bond price at $t+1$, while η_0 is the shadow value of relaxing the participation constraint, η_1 and η_2 are the shadow values for relaxing the

incentive compatibility constraints, and η_3 is the shadow value of relaxing the truth telling constraint. The first order conditions for this problem are:

$$\begin{aligned} v_1(x)^{-1} &= \eta_0 + \eta_1 \left[(\alpha_2/\alpha_1)^{1/(b_t-1)} - g_1(x) \right] - \eta_3 h(x) \\ v_2(x)^{-1} &= \eta_0 + \eta_2 \left[(\alpha_2/\alpha_1)^{1/(b_t-1)} - g_2(x) \right] + \eta_3 \end{aligned}$$

where $h(x) \equiv \varphi_2 f_2(x) / \varphi_1 f_1(x)$ and φ_s denotes the probability the state is s. It is straightforward to show that $\eta_0 = \alpha_2^{1/(b_t-1)}$. If the truth telling constraint is not binding, then $\eta_3 = 0$ and the optimization problem reduces to the pure moral hazard problem solved in Margiotta and Miller (2000). Otherwise $\eta_3 > 0$, and we substitute the first order condition into the incentive compatibility and truth telling constraints, yielding the following system of three equations to solve for the remaining three unknowns η_1, η_2, and η_3:

$$\begin{aligned} & \int_\psi^\infty \frac{1}{\alpha_2^{1/(b_t-1)} - \eta_3 h(x) + \eta_1 \left[(\alpha_2/\alpha_1)^{1/(b_t-1)} - g_1(x) \right]} f_2(x)\, dx \\ = & \int_\psi^\infty \frac{1}{\alpha_2^{1/(b_t-1)} + \eta_3 + \eta_2 \left[(\alpha_2/\alpha_1)^{1/(b_t-1)} - g_2(x) \right]} f_2(x)\, dx \end{aligned}$$

$$\begin{aligned} 0 &= \int_\psi^\infty \frac{g_1(x) - (\alpha_2/\alpha_1)^{1/(b_t-1)}}{\alpha_2^{1/(b_t-1)} - \eta_3 h(x) + \eta_1 (\alpha_2/\alpha_1)^{1/(b_t-1)} - \eta_1 g_1(x)} f_1(x)\, dx \\ &= \int_\psi^\infty \frac{g_2(x) - (\alpha_2/\alpha_1)^{1/(b_t-1)}}{\alpha_2^{1/(b_t-1)} - \eta_3 + \eta_2 (\alpha_2/\alpha_1)^{1/(b_t-1)} - \eta_2 g_2(x)} f_2(x)\, dx \end{aligned}$$

The firms solves similar maximization problems for two of the remaining combinations of effort level, shirking in the first state but working diligently in the second, shirking in the second but not the first, and selects the value maximizing contract.

If there is moral hazard, it is easy to see from the first order conditions that compensation varies with the firm's abnormal returns, exposing the manager to uncertainty. Consequently the firm must pay a risk premium to meet the participation constraint if his compensation is

uncertain and depends on the firm's abnormal returns, because the manager is risk averse. Absent moral hazard, the optimal compensation is a fixed wage award of:

$$w_{j,t+1} = \rho^{-1}(b_t - 1)^{-1} b_{t+1} \log(\alpha_j)$$

for working at effort level j, which just offsets the alternative use of his time. Setting $j = 1$ gives the shirking contract. A rule prohibiting any trading in the firm stock is optimal in this case, and can easily be implemented if the manager's trades are publicly disclosed.

More generally, the compensation schedule should not depend on the manager's private information if, conditional on the manager's effort, the distribution of abnormal returns is independently distributed of the state. To prove this second claim, consider a model where there is only one state, by setting $\varphi_2 = h(x) = 0$. Let $w^*(x)$ denote the optimal contract for the one state model where η_1^* is the associated multiplier for the incentive compatibility constraint. Now suppose $\varphi_2 \neq 0$ but assume $f_1(x) = f_2(x)$ instead. For example the states might be revealed to the manager after he has committed to his effort level but before the end of the period when abnormal returns are realized. In this case he would still personally benefit from insider trading if he was permitted. By assumption, the shirking distributions are the same in both states, meaning $f_1(x) g_1(x) = f_2(x) g_2(x)$, so it now follows that $g_1(x) = g_2(x)$. Hence, upon inspection of the first order conditions and the solution equations for the multipliers, $w^*(x)$ is also the optimal contract for the specialization $f_1(x) = f_2(x)$, and is supported by the multipliers $\eta_1^* = \eta_1 = \eta_2$ with $\eta_3 = 0$. This demonstrates the manager should not be compensated for his hidden information in this case. Unless it is intrinsically tied to the moral hazard problem of motivating them to work diligently, in this model managers should not be allowed to trade on their inside information.

Finally if $\eta_3 > 0$, meaning the truth telling constraint is binding, it follows from arguments in Gayle and Miller (2008b), that the manager is paid more on average in the second state than he would be in the pure moral hazard case, and less in the first state. Intuitively

shareholders must bribe the manager to truthfully reveal the second state when they cannot observe it directly, in order to incentivize his effort. However this gain is offset by a penalty paid in the first state, because competition amongst executives for the position before the state is revealed equalizes their expected utility to the level attained in the pure moral hazard case. The optimal contract exposes managers to uncertainty from two sources, namely the state s, and conditional on the state, the draw of the abnormal return x. Thus the additional uncertainty increases the expected cost of the compensation package to shareholders for the higher risk premium.

IV. Estimating the Costs of Moral Hazard

Our model in Section 3 demonstrates the manager does not profit from an increase in the value of the firm if he signs an optimal contract unless there is a moral hazard problem. As we remarked in the introduction, the disclosure rules of the SEC make it relatively easy for boards to write contracts with managers that prohibit any trading in the firm's securities. Yet the reduced form empirical evidence reported in Section 2 shows that managers benefit significantly from their firm's good fortune. Although linking pay to performance can be rationalized within theoretical models of optimal contracting with moral hazard and hidden actions, the practical relevance of moral hazard to managerial compensation is ultimately an empirical phenomenon, the nub of the third question in the second paragraph of the introduction. Given the risks that insider trading pose for shareholders, is moral hazard a sufficiently important economic factor for firms to incentivize managers?

In this section we estimate the structural parameters of a pure moral hazard model, and compute the costs and benefits to shareholders from incentivizing managers. Our empirical analysis applies nonlinear estimation techniques for parametric models of optimal contracting developed in Margiotta and Miller (2000) and Gayle and Miller (2008a). A companion paper, Gayle and Miller (2008b), analyzes identification, testing and estimation in nonparametric optimal contracting models where there is both moral hazard and private information. Here

we apply parametric methods to a broad industry spectrum of publicly traded firms, thus extending the earlier applications of Margiotta and Miller (2000) and Gayle and Miller (2008a) on smaller data sets to a few select industries. Our new estimates corroborate the earlier work, by showing the estimated losses firms would incur by paying executives a constant wage are much greater than the relatively small amount we estimate shareholders currently pay to incentivize them.

A. Parameterizing the model

There are five parameters to account for systematic differences in executive compensation. They are the probability distribution of abnormal returns conditional on working, $f(x)$, the probability distribution of abnormal returns conditional on shirking, $f(x)\,g(x)$, the risk aversion parameter, ρ, the nonpecuniary benefit from shirking versus working, captured by parameter ratio α_2/α_1, and the nonpecuniary benefit of working versus retiring or accepting employment outside the firm, α_2.

Our empirical analysis assumes $f(x)$ and $f(x)\,g(x)$ are truncated normal with support bounded below by ψ, where (μ_2, σ^2) and (μ_1, σ^2) respectively denote the mean and variance of the parent normal distributions. Thus:

$$f(x) = \frac{\exp\left[-\left(x^2 - 2x\mu_2 + {\mu_2}^2\right)/2\sigma^2\right]}{\Phi\left[\left(\mu_2 - \psi\right)/\sigma\right]\sigma\sqrt{2\pi}}$$

and:

$$g(x) = \frac{\Phi\left[\left(\mu_2 - \psi\right)/\sigma\right]}{\Phi\left[\left(\mu_1 - \psi\right)/\sigma\right]} \exp\left[\left({\mu_2}^2 - {\mu_1}^2 - 2x\mu_2 + 2x\mu_1\right)/2\sigma^2\right]$$

where Φ is the standard normal distribution function. As indicated in the previous section, we cannot reject the null hypothesis of restricting the mean of abnormal returns to zero conditional on working in the data, a restriction we impose in the estimation of the parameter μ_2. This leaves the truncation point ψ, the mean of the parent normal distribution under shirking μ_1, the common variance of the parent normal σ, the risk aversion parameter ρ, the

ratio of nonpecuniary benefits from working to shirking α_2/α_1, and the ratio of nonpecuniary benefits from working to quitting α_2/α_0, to estimate.

The parameters of the distribution of returns are estimated separately for each sector, whereas the taste parameters α_2/α_1 and α_2 are specified as mappings of executive rank. To accommodate other factors that might affect compensation not included in the model we assumed that total compensation, denoted $\widetilde{w}_t$, is the sum of optimal contract compensation w_t plus an independently distributed disturbance term ε_t, assumed orthogonal to the other variables of interest.

B. Parameter Estimates

Table 6 presents the estimates of ψ_i for $i \in \{1, 2, \ldots, 10\}$, the minimal abnormal return defining the lower support point of the truncated normal distribution in the i^{th} sector. The estimators are the minimum difference of the firm return and the market return across all observations in the sector. Because the estimators converge faster than the square root of sample size, so their standard errors (not reported) have no impact on the asymptotic properties of the other parameter estimates.

Loosely speaking, the reported values represent the abnormal return that trigger delisting from the exchange. Our estimates suggest the points at which creditors instigate bankruptcy proceedings, are bought by private investors, or are amalgamated, differ by sector, but are dispersed around the value where the equity value of the firm is close to zero. Since the difference between the firm's financial returns and the return on the market is identically the abnormal return it follows that if the return on a diversified portfolio was r, then an abnormal return of $-r$ would reduce shareholder value to zero, and we previously noted the return on the market portfolio over this period was 1.089.

Table 6

Parameter Estimates of the Diligent-Returns Distribution. Truncation Points

Parameter	Sector	Estimate
ψ_1	Energy	-0.8198
ψ_2	Materials	-0.9812
ψ_3	Industrials	-2.1423
ψ_4	Consumer Discretionary	-1.4905
ψ_5	Consumer Staples	-1.0323
ψ_6	Health Care	-1.0301
ψ_7	Financial	-1.0184
ψ_8	Information Technology	-1.1362
ψ_9	Telecommunication Services	-0.8911
ψ_{10}	Utilities	-0.8097

Note: There is no standard errors because these are super consistent estimates.

In Tables 7 and 8 we report estimates of the three remaining parameters $(\mu_{1i}, \mu_{2i}, \sigma_i)$ that define the truncated normal distribution for each sector $i \in \{1, 2, \ldots, 10\}$. As indicated in the previous section, we cannot reject the null hypothesis of restricting the mean of abnormal returns to zero conditional on working in the data, a restriction we impose in the estimation equations. This explains why μ_{2i} is, without exception, negative and significant.

Table 7

Parameter Estimates of the Diligent-Return Distribution: Means and Standard Deviations of Parent Distribution

Parameter		Sector	Estimates	Standard Error
Standard Deviation	σ_1	Energy	0.898	0.032
	σ_2	Materials	0.333	0.005
	σ_3	Industrials	1.743	0.022
	σ_4	Consumer Discretionary	0.626	0.006
	σ_5	Consumer Staples	0.420	0.008
	σ_6	Health care	42.815	0.775
	σ_7	Financial	0.373	0.004
	σ_8	Information Technology	1.849	0.069
	σ_9	Telecommunication Services	0.579	0.029
	σ_{10}	Utilities	0.289	0.004
Mean	μ_{21}	Energy	-0.5591	0.0592
	μ_{22}	Materials	-0.0017	0.0003
	μ_{23}	Industrials	-0.5652	0.02452
	μ_{24}	Consumer Discretionary	-0.0158	0.0011
	μ_{25}	Consumer Staples	-0.0087	0.0012
	μ_{26}	Health Care	-1608.1984	29.0809
	μ_{27}	Financial	-0.0037	0.0004
	μ_{28}	Information Technology	-2.2483	0.2108
	μ_{29}	Telecommunication Services	-0.0989	0.0207
	μ_{210}	Utilities	-0.0024	0.0003

Table 8

Parameter Estimates of the Shirking-Returns Distribution

Parameter	Description	Sector	Estimate	Standard Error
μ_{11}	Mean return from	Energy	-0.7591	0.0592
μ_{12}	shirking	Materials	-0.037	0.0033
μ_{13}		Industrials	-0.6652	0.0352
μ_{14}		Consumer Discretionary	-0.0458	0.0211
μ_{15}		Consumer Staples	-0.027	0.0312
μ_{16}		Health Care	-1901.19	40.02
μ_{17}		Financial	-0.0097	0.0024
μ_{18}		Information Technology	-4.433	0.4108
μ_{19}		Telecommunication	-0.2989	0.0307
μ_{110}		Utilities	-0.0324	0.0083

The theory predicts that the support for abnormal returns distribution conditional on shirking is contained in the support for abnormal returns conditional on working diligently. Otherwise a first best contract could be achieved by paying the manager a fixed wage supplemented with a sufficiently high fine whenever abnormal returns stray into the region outside the latter, obviating the need for incentive pay that at best can produce a second best contract. In our empirical specification, the two supports are the same, and they share a common sector specific parameter, σ_i, differing only in the mean of the parent distributions, μ_{1i} and μ_{2i}. Thus the mean of the truncated distribution for shirking is less than the mean of the truncated distribution for working diligently if and only if $\mu_{1i} < \mu_{2i}$. Our estimates confirm this is the case in every sector, as Table 8 shows.

Abnormal returns in the health sector behave very differently than the others; although its lower truncation point is the same order of magnitude as in the other sectors, a very low μ_{26} coupled with a very high σ_6 imply the probability density function for abnormal

returns in that sector is estimated to be very flat so that it can capture some high returns that occurred in some firm/years. Similarly our estimate of μ_{1i} is orders of magnitude lower than the counterparts for the other sectors, signifying an even flatter density for the shirking distribution.

The preference parameter estimates are presented in Table 9. Our estimate of ρ, the risk aversion parameter, implies utility is concave increasing as required by the model, and lies between results in reported in Margiotta and Miller (2000) and Gayle and Miller(2008a). It implies that a manager would be indifferent between accepting a lottery offering even odds of winning or losing one million dollars versus losing an amount of $\$103,259$ for sure. Our estimate of ξ, the parameter explaining variation not captured by the model, is similar too. (More precisely, the variance of the measurement error is $2b_{t+1}\rho^{-2}\xi$.)

We estimated α_2, the parameter determining competitive opportunities in the labor market for executives, and α_2/α_1, a measure of nonpecuniary benefits from shirking versus working diligently, by executive position for the top six ranked executives. All our estimates of α_2 are greater than one numerically, but only in the upper ranks is the null hypothesis that $\alpha_2 = 1$ rejected in favor of the one sided alternative $\alpha_2 > 1$. Recalling that the exponential utility function is negative, these results are weak evidence that the nonpecuniary benefits of the job relative to outside opportunities in the labor market, decline with promotion. It appears that financial remuneration, rather than power, prestige or perks, is necessary to motivate executives to climb the corporate ladder. Our estimates of α_2/α_1 show that chief executive officers would benefit significantly from taking actions that are not in the shareholder's interests if they are not incentivized, whereas the opportunities afforded to lower ranked executives if their pay is not tied to performance are more limited, presumably because discretion about work activities and job duties increase with rank while the degree of supervision declines. These issues are investigated more thoroughly in Gayle, Golan and Miller (2008), who estimate the dynamic life cycle aspects of executive promotion and turnover.

Table 9

Parameter Estimates of the Manager's Utility Function

Parameter	Description	Executive Rank	Estimate	Standard Error
ρ	Risk aversion		0.208	0.102
ξ	Measurement error		2.03	0.505
α_2	Preference for	CEO	1.292	0.0162
	diligence relative	2nd ranked	1.523	0.126
	to quitting firm	3rd ranked	1.420	0.118
		4th ranked	1.48	0.375
		5th ranked	1.373	0.504
		6th ranked	1.849	0.969
α_2/α_1	Preference for	CEO	1.356	0.129
	diligence relative	2nd ranked	1.034	0.034
	to shirking	3rd ranked	1.012	0.045
		4th ranked	1.023	0.078
		5th ranked	1.01	0.678
		6th ranked	0.987	0.567

C. The Costs of Moral Hazard

We characterize the importance of moral hazard three ways, the gross loss shareholders would incur before accounting for managerial compensation from the manager tending his own interests, the benefits accruing to the manager from tending his own interests instead of his shareholders, and how much the shareholders are willing to pay to eliminate the problem of moral hazard altogether.

Table 10

Gross Loss to Firms from not Controlling Moral Hazard

in millions of US$ (2000)

Sector	Estimate of τ_1
Energy	1,290
Materials	1,468
Industrials	1,679
Consumer Discretionary	1,235
Consumer Staples	987
Health Care	2,877
Financial	1,568
Information Technology	1,457
Telecommunication	1,078
Utilities	569

The first measure, denoted τ_1, is the expected gross output loss to the firm switching from the distribution of abnormal returns for diligent work to the distribution for shirking, which has probability density function $f(x)\,g(x)$. Stated differently, τ_1 is the difference between the expected output to the plant from the manager pursuing the firm's goals versus his own, before netting out expected managerial compensation:

$$\tau_1 = -V \int xg(x)\, f(x)\, dx \equiv -VE\left[xg(x)\right]$$

where V is the equity value of the firm, this formula exploiting the identity that the expected value of abnormal returns is zero when the manager pursues the interests of the firm. Table 10 displays the estimated average over all firms in each sector for withdrawing the incentives for the managers to work diligently. Comparing these numbers with the size of firms reported in Table 3, we find that the value of equity would decline precipitously if managers were not

incentivized to align their personal objectives with those of the firms they manage. This result essentially replicates the findings of Margiotta and Miller (2000) and Gayle and Miller (2008a) for a much smaller select group of narrowly defined industries (aerospace, chemicals and electronics).

The second measure, τ_2, is the nonpecuniary benefits to management from ongoing shirking, that is successive managers pursuing their own goals within the firm each period. Suppressing the time subscript, and supposing that the bond price b is constant, let w_2 denote the manager's reservation wage to work under perfect monitoring or if there were no moral hazard problem, that is his certainty equivalent of the current compensation package, and let w_1 denote the manager's reservation wage to shirk. Then the present value of the compensating differential for these two activities, can be expressed as the difference $\tau_2 \equiv b\left(w_2 - w_1\right)$.

Our estimates of τ_2 for the top executive position is $\$24,690,192$, for the second in command $\$4,460,774$, tiny compared to the expected losses a firm would incur; our model predicts there are enormous gains from having managers act in the interests of shareholders. (Estimates of the lower ranked executives are considerably lower than for the second in command.) From the manager's perspective, however, the annuity implied by τ_2 is quite substantial, and for a sizeable proportion of the sample population exceeds annual compensation. Of course if a manager decided to shirk to receive these sizable nonpecuniary benefits then his expected compensation would fall drastically because his inside wealth would lose much of its value.

Finally we estimated the maximal amount shareholders are willing pay to eliminate the moral hazard problem, the value of a perfect monitor, τ_3. Absent moral hazard, the firm would pay the manager the fixed wage w_2, instead of according to the compensation schedule $w\left(x\right)$. Another way of expressing τ_3 is the equilibrium risk premium paid to an executive for taking a job that offers an uncertain income. The firms' willingness to pay for eliminating the moral hazard problem for one period is therefore $\tau_3 \equiv E\left[w\left(x\right)\right] - w_2$. We computed this

measure for the CEO and second highest paid executive for each sector.[4]

Table 11

Cost of Moral Hazard in US$ (2000)

Sector	Executive	Estimate of τ_3
Energy	CEO	10,450,320
	2nd ranked	1,345,098
Materials	CEO	11,450,450
	2nd ranked	1,745,067
Industrials	CEO	14,670,350
	2nd ranked	1,675,067
Consumer Discretionary	CEO	8,210,950
	2nd ranked	3,245,067
Consumer Staples	CEO	4,210,950
	2nd ranked	545,068
Health care	CEO	30,410,580
	2nd ranked	10,450,000
Information Technology	CEO	12,410,580
	2nd ranked	4,550,134
Telecommunication	CEO	15,670,892
	2nd ranked	4,550,134
Utilities	CEO	6,590,872
	2nd ranked	450,674

Confirming our previous work, the estimates in Table 11 demonstrate that the risk premium paid to executives is a very important part of their pay package. Elsewhere we have argued, in Gayle and Miller (2008b), that changes in this component are largely responsible

[4]Note that since the probability distribution of abnormal returns conditional on effort level varies with the sector, the cost of moral hazard is sector dependent too.

for expected compensation and its volatility, increasing faster than real wages over the last 60 years. Here we simply add that there is notable variation between the costs of moral hazard across the sectors, with the health care sector registering as an outlier worthy of special attention in a future study.

V. Conclusion

The disclosure rules of the SEC make it relatively easy for boards to write contracts with managers that prohibit any trading in the firm's securities. Yet our empirical evidence shows that managers significantly benefit from their firm's good fortune. Consistent with previous work, in this area we find that managers exploit insider information about the profitability of their own firm for direct personal gain. But this is not sufficient to prove that executive compensation contracts are deficient. Our theoretical framework demonstrates that in an optimal contract the manager should not profit from changes in the value of the firm unless there is a moral hazard problem. This feature might explain the paradox of inside information and performance pay. Optimal contracting in models of generalized moral hazard, with both private information and hidden actions, reward managers for truthfully revealing the state of the firm. Shareholders permit compensation schemes that correlate firm performance with executive pay because the profitability of the firm depends on how managers assess their own accomplishments and firm's prospects, as well as what managers do, which is organizing human resources in creative ways that add value to their firm. Rewarding managers for revealing hidden information about the firm's profitability helps the board set contracts that incentivize the manager's work activities. If moral hazard is anywhere near as costly as our estimated values, then de-coupling managerial compensation from changes in shareholder wealth, for example by imposing an upper limit on maximal compensation, could be very costly indeed.

Notes

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